

Variance

CS 70, Summer 2019

Lecture 21, 7/30/19

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Two Games ^{fair}

Game 1: Flip a ^{fair} coin 10 times. For each Head, you win 100. For each Tail, you lose 100.

Expected Winnings on Flip i :

F_i

$$\mathbb{E}[F_i] = 100\left(\frac{1}{2}\right) + (-100)\left(\frac{1}{2}\right) = 0$$

Expected Winnings After 10 Flips:

F

$$\mathbb{E}[F] = \sum_{i=1}^{10} \mathbb{E}[F_i] = 0$$

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Two Games ^{fair}

Game 2: Flip a ^{fair} coin 10 times. For each Head, you win 10000. For each Tail, you lose 10000.

Expected Winnings on Flip i :

F_i

$$\mathbb{E}[F_i] = 0 = \frac{1}{2}(10000) + \frac{1}{2}(-10000)$$

Expected Winnings After 10 Flips:

F

$$\mathbb{E}[F] = 0$$

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Q: Which game would you rather play?

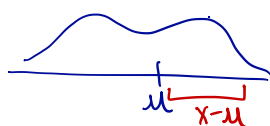
Definition of Variance

The key difference is the **variance**.

Variance is the **expected "distance" to mean**.

Let X be a RV with $\mathbb{E}[X] = \mu$. Then:

$$\text{Var}(X) = \mathbb{E}[(X - \mu)^2]$$



variance is always non-neg

Std Dev:

$$\sigma(X) = \sqrt{\text{var}(X)}$$

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Alternate Definition

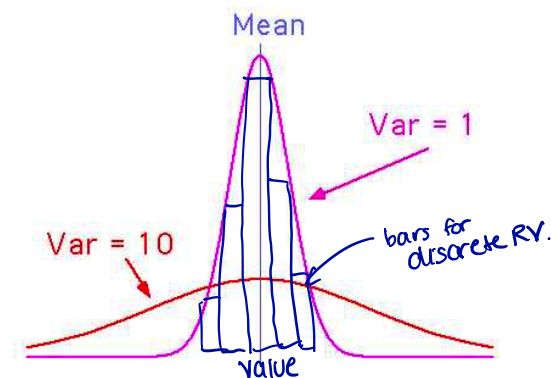
We can use **linearity of expectation** to get an alternate form that is often **easier to apply**.

$$\text{Var}(X) = \mathbb{E}[X^2] - \mu^2$$

$$\begin{aligned} \mathbb{E}[(X - \mu)^2] &= \mathbb{E}[X^2 - 2\mu X + \mu^2] \\ \text{linearity} \rightarrow &= \mathbb{E}[X^2] - 2\mu \mathbb{E}[X] + \mu^2 \\ &= \mathbb{E}[X^2] - \mu^2 \end{aligned}$$

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Variance: A Visual



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Variance of a Bernoulli

Let $X \sim \text{Bernoulli}(p)$.

Then $\mathbb{E}[X] = p$

What is X^2 ? $\mathbb{E}[X^2]$?

$$X^2 = \begin{cases} 1 & \text{wp } p \\ 0 & \text{wp } 1-p \end{cases} \quad \mathbb{E}[X^2] = 1 \cdot p + 0 \cdot (1-p) = p$$

$$\text{Var}[X] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2 = p - p^2 = p(1-p)$$

$X \sim \text{Ber}(p), Y \sim \text{Ber}(1-p)$
 $\text{Var}(X) = \text{Var}(Y)$

Variance of a Dice Roll

What is the variance of a single 6-sided dice roll?

$R = \text{value of a dice roll. } \{1, 2, 3, 4, 5, 6\}$

What is R^2 ?

$$R^2 = \begin{cases} 1 & \text{wp } \frac{1}{6} \\ 4 & \text{wp } \frac{1}{6} \\ 9 & \text{wp } \frac{1}{6} \\ 16 & \text{wp } \frac{1}{6} \\ 25 & \text{wp } \frac{1}{6} \\ 36 & \text{wp } \frac{1}{6} \end{cases}$$

Variance of a Dice Roll

$$\mathbb{E}[R^2] = \frac{1}{6} [1 + 4 + 9 + 16 + 25 + 36]$$

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$$= \frac{1}{6} (91)$$

$$\text{Var}(R) = \mathbb{E}[R^2] - (\mathbb{E}[R])^2 = \frac{91}{6} - \left(\frac{7}{2}\right)^2$$

(Notes.)

Variance of a Geometric

Know the variance; proof optional, but good practice with **manipulating RVs**.

Let $X \sim \text{Geometric}(p)$.

Strategy: Nice expression for $p \cdot \mathbb{E}[X^2]$

$$\mathbb{E}[X^2] = 1 \cdot p + 4(1-p)p + 9(1-p)^2p + \dots$$

$$- [1-p]\mathbb{E}[X^2] = - [1 \cdot (1-p)p + 4(1-p)^2p + \dots]$$

subtract from both sides.

$$p\mathbb{E}[X^2] = 1 \cdot p + 3(1-p)p + 5(1-p)^2p + \dots$$

$$= (2 \cdot p + 4(1-p)p + 6(1-p)^2p + \dots) + (-p - (1-p)p - (1-p)^2p)$$

Variance of a Geometric II

From the distribution of X , we know:

$$\textcircled{1} \sum_{i=1}^{\infty} p[X=i] = p + (1-p)p + (1-p)^2p + \dots = 1$$

From $\mathbb{E}[X]$, we know:

$$\mathbb{E}[X] = 1 \cdot p + 2 \cdot (1-p)p + 3(1-p)^2p + \dots$$

$$\textcircled{2} \quad \frac{1}{p}$$

$$(2 \cdot p + 4(1-p)p + 6(1-p)^2p + \dots) + (-p - (1-p)p - (1-p)^2p)$$

$$p\mathbb{E}[X^2] = 2\mathbb{E}[X] - 1 \quad \text{solve for } \mathbb{E}[X^2]$$

$$\mathbb{E}[X^2] = \frac{2-p}{p^2}$$

Variance of a Geometric III

Recall $\mathbb{E}[X] = \frac{1}{p}$.

$$\text{var}(X) = \mathbb{E}[X^2] - (\mathbb{E}[X])^2 = \frac{2-p}{p^2} - \frac{1}{p^2} = \boxed{\frac{1-p}{p^2}}$$

Variance of a Poisson

Same: **know the variance**; proof optional, but good practice with **functions of RVs**.

Let $X \sim \text{Poisson}(\lambda)$.

Strategy: Compute $\mathbb{E}[X(X-1)]$.

$$\begin{aligned}\mathbb{E}[X(X-1)] &= \sum_{i=0}^{\infty} i(i-1) \cdot \underbrace{\frac{\lambda^i}{i!} e^{-\lambda}}_{\mathbb{P}[X=i]} \\ &= e^{-\lambda} \sum_{i=2}^{\infty} \frac{\lambda^i}{(i-2)!} \quad \text{Taylor series for } e^{\lambda} \\ &= e^{-\lambda} \lambda^2 \sum_{i=2}^{\infty} \frac{\lambda^{i-2}}{(i-2)!} \\ &= \cancel{e^{-\lambda}} (\lambda^2) (\cancel{e^{\lambda}}) = \lambda^2\end{aligned}$$

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Variance of a Poisson II

Use $\mathbb{E}[X(X-1)]$ to compute $\text{Var}(X)$.

$$\begin{aligned}\text{Var}(X) &= \mathbb{E}[X^2] - (\mathbb{E}[X])^2 \\ &= \mathbb{E}[X(X-1)] + \mathbb{E}[X] - (\mathbb{E}[X])^2 \\ &= \underbrace{\lambda^2}_{\text{last slide}} + \underbrace{\lambda}_{\text{yes! random}} - \lambda^2 = \lambda\end{aligned}$$

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Break

Would you rather only wear sweatpants for the rest of your life, or never get to wear sweatpants ever again?

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Properties of Variance I: Scale

Let X be a RV, and let $c \in \mathbb{R}$ be a constant.

Let $\mathbb{E}[X] = \mu$.

$$\text{Var}(cX) = c^2 \cdot \text{Var}(X)$$

$$\begin{aligned}\text{Var}(cX) &= \mathbb{E}[(cX)^2] - (\mathbb{E}[cX])^2 \quad \text{lin.} \\ &= \mathbb{E}[c^2 X^2] - (c \mathbb{E}[X])^2 \\ &= c^2 \mathbb{E}[X^2] - c^2 (\mathbb{E}[X])^2 \\ &= c^2 \text{Var}(X)\end{aligned}$$

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Properties of Variance II: Shift

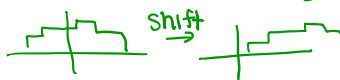
Let X be a RV, and let $c \in \mathbb{R}$ be a constant.

Let $\mathbb{E}[X] = \mu$.

Then, let $\mu' = \mathbb{E}[X + c] = \mu + c$

$$\text{Var}(X + c) = \text{Var}(X)$$

$$\begin{aligned}\text{Var}(X + c) &= \mathbb{E}[(X + c) - \mu']^2 \\ &= \mathbb{E}[(X + c - \mu - c)^2] \\ &= \mathbb{E}[(X - \mu)^2] = \text{Var}(X)\end{aligned}$$



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Example: Shift It!

Consider the following RV:

$$X = \begin{cases} 1 & \text{w.p. } 0.4 \\ 3 & \text{w.p. } 0.2 \\ 5 & \text{w.p. } 0.4 \end{cases}$$

$$= \text{Var}(X-3)$$

What is $\text{Var}(X)$? Shift it!

$$\begin{aligned}X-3 &= \begin{cases} -2 & \text{wp. } 0.4 \\ 0 & \text{wp. } 0.2 \\ 2 & \text{wp. } 0.4 \end{cases} \quad \left| \quad \begin{aligned} (X-3)^2 &= \begin{cases} 4 & \text{wp. } 0.8 \\ 0 & \text{wp. } 0.2 \end{cases} \\ \mathbb{E}[(X-3)^2] &= 4 \cdot 0.8 + 0 \cdot 0.2 = 3.2 \\ \mathbb{E}[X-3] &= 0 \\ \text{Var}(X-3) &= 3.2 \end{aligned}\end{aligned}$$

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Sum of Independent RVs

Let X_1, \dots, X_n be **independent** RVs. Then:

$$\text{Var}(X_1 + \dots + X_n) = \text{Var}(X_1) + \dots + \text{Var}(X_n)$$

Proof: Tomorrow!

Today: Focus on applications.

Variance of a Binomial

Let $X \sim \text{Bin}(n, p)$. Then,

$$X = X_1 + X_2 + \dots + X_n$$

Here, $X_i \sim \text{Ber}(p)$ X_i iid.

$$\text{Var}(X) = \text{Var}(X_1) + \text{Var}(X_2) + \dots + \text{Var}(X_n)$$

$$= n \cdot \text{Var}(X_1)$$

$$= \boxed{n p(1-p)}$$

$X \sim \text{Bin}(n, p)$
 $Y \sim \text{Bin}(n, 1-p)$
 $\text{Var}(X) = \text{Var}(Y)$

Sum of Dependent RVs

Main strategy: **linearity of expectation** and **indicator variables**

Useful Fact:

$$(X_1 + X_2 + \dots + X_n)^2 =$$

$$(X_1 + X_2 + \dots + X_n)(X_1 + X_2 + \dots + X_n)$$

$$= (X_1^2 + X_2^2 + \dots + X_n^2) + (X_1 X_2 + X_1 X_3 + \dots + X_{n-1} X_n)$$

$$= \underbrace{\sum_{i=1}^n X_i^2}_n + \underbrace{\sum_{i \neq j} X_i X_j}_{n(n-1)}$$

alternate:
 $2 \sum_{i < j} X_i X_j$

HW Mixups (Fixed Points)

(In notes.) n students hand in HW. I mix up their HW randomly and return it, so that every possible mixup is **equally likely**.

Let $S = \#$ of students who get their own HW.

Last time: defined $S_i =$ indicator for student i getting own HW.

$$S_i \sim \text{Ber}\left(\frac{1}{n}\right)$$

Using **linearity of expectation**:

$$\mathbb{E}[S] = \mathbb{E}[S_1 + S_2 + \dots + S_n] = \mathbb{E}[S_1] + \mathbb{E}[S_2] + \dots + \mathbb{E}[S_n]$$

$$= \frac{1}{n}(n) \text{ (1)}$$

HW Mixups II

Using our **useful fact**:

$$\mathbb{E}[S^2] = \mathbb{E}[(S_1 + S_2 + \dots + S_n)^2]$$

$$= \mathbb{E}\left[\sum_{i=1}^n S_i^2 + \sum_{i \neq j} S_i S_j\right]$$

linearity:

$$= \mathbb{E}\left[\sum_{i=1}^n S_i^2\right] + \mathbb{E}\left[\sum_{i \neq j} S_i S_j\right]$$

$$= n \cdot \mathbb{E}[S_1^2] + n(n-1) \mathbb{E}[S_1 S_2]$$

? ?

HW Mixups III

What is S_i^2 ? $\mathbb{E}[S_i^2]$?

$$S_i^2 = \begin{cases} 1 & \text{wp } \frac{1}{n} \\ 0 & \text{wp } 1 - \frac{1}{n} \end{cases} \quad \mathbb{E}[S_i^2] = \frac{1}{n}$$

For $i \neq j$, what is $S_i S_j$? $\mathbb{E}[S_i S_j]$?

$$S_i S_j = \begin{cases} 1 & \Rightarrow \text{both } S_i, S_j = 1 \\ 0 & \Rightarrow \text{both student } i \text{ and } j \text{ get own HW.} \end{cases} \quad \mathbb{E}[S_i S_j] = \frac{1}{n} \left(\frac{1}{n-1}\right)$$

HW Mixups IV

Put it all together to compute $\text{Var}(X)$.

$$\begin{aligned}\text{Var}(X) &= \underbrace{\mathbb{E}[X^2]} - (\mathbb{E}[X])^2 \\ &= n\mathbb{E}[S_1^2] + n(n-1)\mathbb{E}[S_1S_2] - (1)^2 \\ &= n\left(\frac{1}{n}\right) + n(n-1)\frac{1}{n(n-1)} - 1 \\ &= 1\end{aligned}$$

Summary

Today:

- ▶ Variance measures how far you **deviate from mean**
- ▶ Variance is additive for **independent RVs**; proof to come tomorrow
- ▶ Use **linearity of expectation** and **indicator variables**