

Intro to Markov Chains

CS 70, Summer 2019

Lecture 26, 8/7/19

Applications of Markov Chains

- ▶ Models systems of **states** and **transitions**
- ▶ PageRank – Google's search algorithm.
States are webpages, transitions are links.
- ▶ Tons of applications outside of CS: statistical physics, speech recognition, bioinformatics, sabermetrics...

↳ baseball statistics!

Markov Chain Definition

Three key components (and one assumption):

- ▶ Set \mathcal{S} of **states**

Think of these as **vertices of a graph**

- ▶ Transition probabilities. $P[i \rightarrow j]$

Think of these as **directed edges in a graph**

Transitions **out of a node** should sum to **1**

- ▶ Initial **distribution** $\mu^{(0)}$.

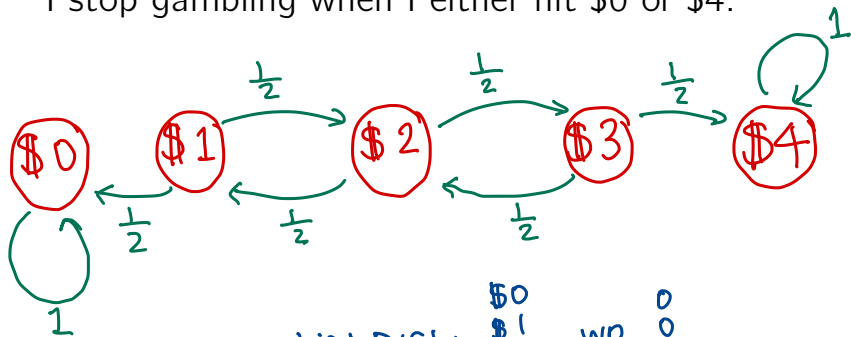
Gives the probability that we start at a state.

- ▶ Memorylessness (aka **Markov property**)

Example: Gambling

I start with \$2. If I guess a coin flip correctly, I get \$1, and if I am incorrect, I lose \$1.

I stop gambling when I either hit \$0 or \$4.



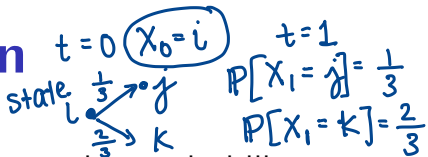
Initial Dist:

\$0
\$1
\$2
\$3
\$4

wp

0
0
1
0
0

Traversing the Chain



X_0 is the initial state.

Choose transitions according to its probability.

X_i is the state you're on at time i . X_i is a **RV**.

Markov Property: ("Memory less")

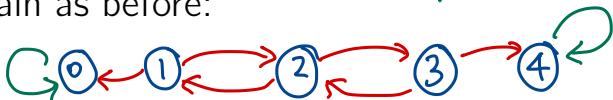
Only the **current state** matters for the next.

"Knowing the entire history of the chain is equivalent to just knowing the current state."

$$\begin{aligned} P[X_{n+1} = S_{n+1} \mid X_0 = S_0, X_1 = S_1, \dots, X_n = S_n] \\ = P[X_{n+1} = S_{n+1} \mid X_n = S_n] \end{aligned}$$

Gambling II

Same chain as before:



What is $\mathbb{P}[X_1 = 3 | X_0 = 2]$? $\frac{1}{2}$

What is $\mathbb{P}[X_{100} = 3 | X_{99} = 2, X_0 = 2]$? $\rightarrow \mathbb{P}[X_{100} = 3 | X_{99} = 2]$
 $= \mathbb{P}[X_1 = 3 | X_0 = 2]$

What is $\mathbb{P}[X_1 = 3, X_2 = 2, X_3 = 3, X_4 = 4]$? $= \frac{1}{2}$

$$\frac{1}{2} \times \frac{1}{2} \times \frac{1}{2} \times \frac{1}{2} = \frac{1}{16}$$

$$\mathbb{P}[X_1 = 3 | X_0 = 2] \downarrow$$
$$\mathbb{P}[X_2 = 2 | X_1 = 3]$$

$X_0 = 2$
↑
same initial setup

Gambling II

What is $\mathbb{P}[X_4 = 4]$? Same setup $X_0 = 2$

X_0 X_1 X_2 X_3 $X_4 = 4$

Path 1: 2 3 4 4 4

Path 2: 2 3 2 3 4

Path 3: 2 1 2 3 4

$$\mathbb{P}[X_4 = 4] = \mathbb{P}[\text{Path 1}] + \mathbb{P}[\text{Path 2}] + \mathbb{P}[\text{Path 3}]$$

↑
Good idea, but it doesn't scale.

The Transition Matrix

Calculations are easier to do when we stick the transition probabilities in a matrix.

Transition matrix P .

The (i, j) entry is $\mathbb{P}[X_1 = j | X_0 = i]$, or the transition from i to j .

$P = \begin{bmatrix} \mathbb{P}[1 \rightarrow 1] & \mathbb{P}[1 \rightarrow 2] & \dots \\ \mathbb{P}[2 \rightarrow 1] & \dots & \dots \\ \vdots & \dots & \dots \\ \vdots & \dots & \dots \end{bmatrix}$

Handwritten annotations:
- Above the matrix: $\mathbb{P}[i \rightarrow j]$
- To the right: \leftarrow col index = States
- To the left: \downarrow column.
- To the left: \uparrow row
- To the left: \rightarrow row index = States

Sanity check:

row sum: 1

col sum: non-neg.,
otherwise no restriction

entries: non-neg
 ≤ 1 .

The Distribution Vectors

So far: saw initial distribution $\mu^{(0)}$.

Can represent it as a **row vector**:

$$\begin{array}{cccc} & [P[X_0=1] & P[X_0=2] & \dots \\ \text{index} \rightarrow & 1 & 2 & 3 \dots \\ = \text{states} & & & \end{array} \quad \left. \vphantom{\begin{array}{cccc} & [P[X_0=1] & P[X_0=2] & \dots \\ \text{index} \rightarrow & 1 & 2 & 3 \dots \\ = \text{states} & & & \end{array}} \right\} \begin{array}{l} \text{entries} \\ \text{sum to 1.} \end{array}$$

We can also define a **distribution at time n** :

$$\begin{array}{cccc} & [P[X_n=1] & P[X_n=2] & \dots \\ \text{index} \rightarrow & 1 & 2 & 3 \dots \\ = \text{states} & & & \end{array} \quad \left. \vphantom{\begin{array}{cccc} & [P[X_n=1] & P[X_n=2] & \dots \\ \text{index} \rightarrow & 1 & 2 & 3 \dots \\ = \text{states} & & & \end{array}} \right\} \begin{array}{l} \text{call this} \\ \mu^{(n)} \end{array}$$

Distribution at Time 1

$$[P[X_0=1] \quad P[X_0=2] \quad \dots]$$

We'll prove that $\mu^{(0)} P = \mu^{(1)}$

$$1 \times n \quad n \times n \quad 1 \times n$$

$$\begin{bmatrix} P[1 \rightarrow i] \\ P[2 \rightarrow i] \\ \vdots \end{bmatrix}$$

$$P[X_1=i] = i^{\text{th}} \text{ entry of } \mu^{(1)}$$

$$= \mu^{(0)} \times i^{\text{th}} \text{ column of } P$$

$$= P[X_0=1] \times P[X_1=i | X_0=1] + P[X_0=2] \times P[X_1=i | X_0=2] + \dots$$

casework on X_0

$$= P[X_1=i] \leftarrow \text{Total Prob. Rule.}$$

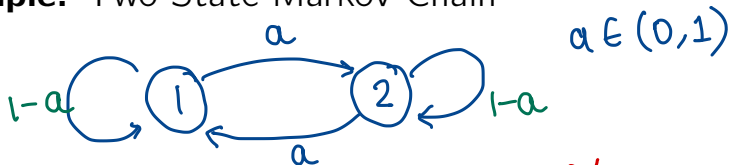
If we know $\mu^{(1)}$, how do we get $\mu^{(2)}$?

$$\mu^{(2)} = \mu^{(1)} P = \mu^{(0)} P^2$$

Distribution at Time n

In general: $\mu^{(n)} = P^n \mu^{(0)}$. (Proof optional.)

Example: Two State Markov Chain



$$P = \begin{bmatrix} 1-a & a \\ a & 1-a \end{bmatrix}$$

$$(Notes:) P^n = \begin{bmatrix} \underbrace{\frac{1}{2} + \frac{1}{2}(1-2a)^n}_{c'} & \frac{1}{2} - \frac{1}{2}(1-2a)^n \\ \frac{1}{2} - \frac{1}{2}(1-2a)^n & \underbrace{\frac{1}{2} + \frac{1}{2}(1-2a)^n}_{c} \end{bmatrix}$$

prove using induction

Aside: $n \rightarrow \infty$

For the two state Markov chain, as $n \rightarrow \infty$,

$$P^n \rightarrow \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix}$$

No matter what $\mu^{(0)}$ is:

$$[p \quad 1-p]$$

$$[p \quad 1-p] \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix} = \left[\frac{1}{2} \quad \frac{1}{2} \right]$$

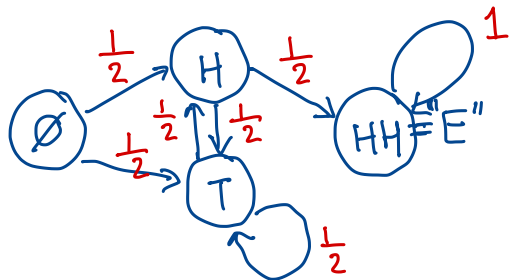
Tomorrow: we'll study this in greater detail!

Break

What's the weirdest thing you've ever eaten?

First Step Analysis: Two Heads

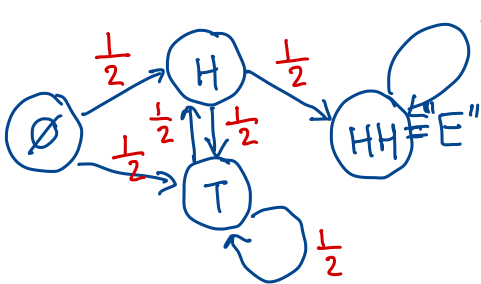
I repeatedly flip a coin, and stop when I get two heads in a row. What is the expected number of flips I need before stopping?



First Step Analysis: Two Heads

For state S , let $\tau(S)$ be the expected time to two heads, starting from state S . 4 variables: $\tau(\emptyset)$

Analyze a **single transition** out of each state to get the **first step equations**: $\tau(H)$
 $\tau(T)$
 $\tau(E)$



$$1 \ \emptyset: \tau(\emptyset) = 1 + \frac{1}{2}\tau(H) + \frac{1}{2}\tau(T)$$

$$T: \tau(T) = 1 + \frac{1}{2}\tau(H) + \frac{1}{2}\tau(T)$$

$$H: \tau(H) = 1 + \frac{1}{2}\tau(T) + \frac{1}{2}\tau(E)$$

$$HH: \tau(E) = 0$$

Goal: $\tau(\emptyset)$. Notes.

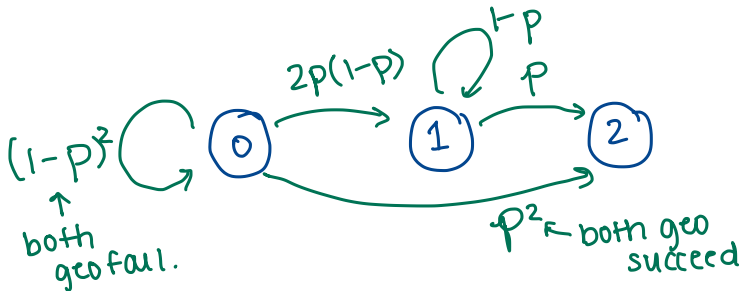
Max of Two Geometrics

Let $X, Y \sim \text{Geometric}(p)$. X, Y are **independent**.

Say X, Y model time until a success.

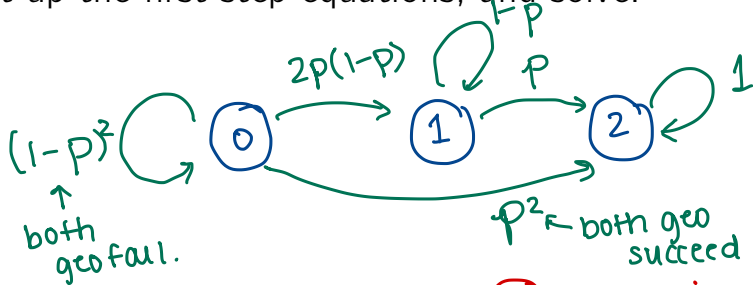
$\max(X, Y)$ is the first time that both X, Y have succeeded at least once. What is $\mathbb{E}[\max(X, Y)]$?

states = # successes



Max of Two Geometrics

Set up the first step equations, and solve:



$\tau(i)$ = expected time until (2) from i

$$0: \tau(0) = 1 + (1-p)^2 \tau(0) + 2p(1-p) \tau(1) + p^2 \tau(2)$$

$$1: \tau(1) = 1 + (1-p) \tau(1) + p \tau(2)$$

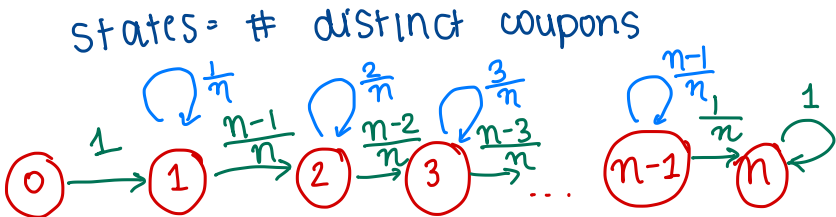
$$2: \tau(2) = 0$$

Exercise: Finish up.

Coupon Collector: A Markov Chain?

Can we reformulate Coupon Collector (with n distinct coupons) as a Markov chain?

How do we recover the expected number of coupons needed to get all n distinct ones?

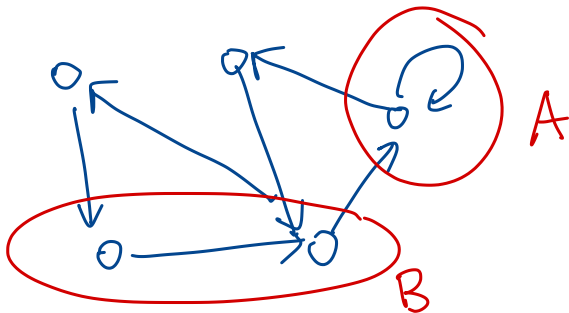


Let $\tau(i)$ = expected time to "n" from "i"
Goal: $\tau(0)$.

Probability of \mathcal{A} Before \mathcal{B}

Let \mathcal{A} and \mathcal{B} be two **disjoint** subsets of the states \mathcal{S} of a Markov chain.

Let $\alpha(i)$ be the probability that we enter \mathcal{A} before entering \mathcal{B} , if we start at state i .



Probability of A Before B

Can also run first step analysis!

$$\text{If } i \in A: \alpha(i) = 1$$

Already in A !

$$i \in B: \alpha(i) = 0$$

Impossible to get to A before B .

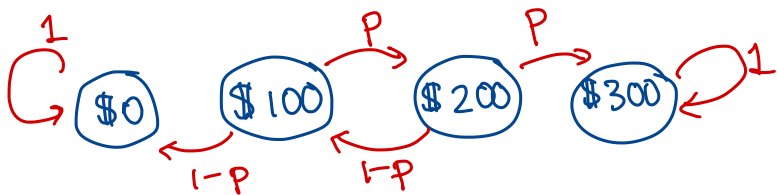
$$\text{else: } \alpha(i) = \sum_{\substack{\text{neighbors } j \\ \text{of } i}} P[i \rightarrow j] \alpha(j)$$

Case work based on taking 1 step from i .

Gambling III

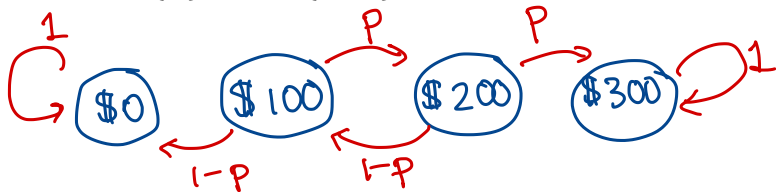
I start with \$100. In each round, I win \$100 with probability p and lose \$100 with probability $(1 - p)$. I end when I either have \$0 or \$300.

What is the probability I end the game with \$300?



Gambling III

Let $\mathcal{A} = \{0\}$, $\mathcal{B} = \{300\}$. First step equations:



$\alpha(i) = \mathbb{P}[\mathcal{B} \text{ before } \mathcal{A} \mid \text{at state } i]$

$$\text{\$ } 0: \alpha(0) = 0$$

$$\text{\$ } 100: \alpha(100) = p\alpha(200) + (1-p)\alpha(0)$$

$$\text{\$ } 200: \alpha(200) = p\alpha(300) + (1-p)\alpha(100)$$

$$\text{\$ } 300: \alpha(300) = 1 \quad \Rightarrow \quad \alpha(100) = p\alpha(200)$$

$$\Rightarrow \alpha(100) = \frac{p^2 \alpha(200)}{1-p(1-p)}$$

Summary

- ▶ Markov chains let you model real world problems with **states** and **transition probabilities**
- ▶ The **Markov property** tells you that where you go next only depends on the **current state**, not on any previous history.
- ▶ The first step analysis is a simple way of analyzing expected hitting times and probabilities of hitting certain states before others.